

# Sampling from a Probability Distribution

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We introduce quantiles, and show that sampling a quantile uniformly at random, and then inverting a distribution  $F$  at that sample value, is equivalent to sampling from  $F$  itself.

## 1 Quantiles

A **quantile**  $q \in [0, 1]$  is the relative strength of a value  $v \in T$ :

$$q(v) = 1 - F(v).$$

Likewise,

$$F(v) = 1 - q(v).$$

The value  $F(v)$  is the probability that the value of a random draw from the distribution  $F$  is less than or equal to  $v$ . Accordingly, the quantile  $v$  is the probability that the value of a random draw is greater than  $v$ . Thus, lower quantiles correspond to higher types/values, and higher quantiles correspond to lower types/values.

To map a quantile back to a value, we invert  $F$  at  $1 - q$ :

$$v(q) = F^{-1}(1 - q).$$

Perhaps the most famous quantile is the **median**, which is the value of a CDF  $F$  inverted at  $1/2$ . The 10th% quantile corresponds to the value at the 90th% percentile, meaning the value below which 90% percent of the data lie. (And likewise, for the  $k$ th quantile and the  $100 - k$ th percentile.)

**N.B.** Quantiles are more often defined as  $F(v)$ , rather than  $1 - F(v)$ . Our choice is natural in the context of this lecture, whose subject is posted-price revenue maximization, because the probability of a sale is the probability that a buyer's value (i.e., a draw from  $F$ ) exceeds the posted price.

## 2 Probability Integral Transform

Let  $X$  be a random variable with arbitrary CDF  $F$ . Now define the random variable  $Y = F(X)$ . Note that the range of  $Y$  is  $[0, 1]$ . One interesting question is, how is  $Y$  distributed? That is, *what is the CDF of an arbitrary CDF*? The answer to this question is straightforward to derive, but surprising nonetheless.

For  $0 \leq y \leq 1$ ,

$$\begin{aligned}\Pr(Y \leq y) &= \Pr(F(X) \leq y) \\ &= \Pr(X \leq F^{-1}(y)) \\ &= F(F^{-1}(y)) \\ &= y.\end{aligned}$$

Since  $\Pr(Y \leq y) = y$ ,  $Y = F(X)$  is uniformly distributed on  $[0, 1]$ ! That is, *the CDF of an arbitrary CDF is uniform*. Finally, since a quantile associated with a value is 1 less the CDF at that value, it follows that *quantiles are likewise uniformly distributed*.

### 3 An Aside: Inverse Transform Sampling

Let's apply this same logic to the random variable  $X$ , defined as the inverse of the uniform distribution for some cdf  $F$  (i.e.,  $X = F^{-1}(U)$ ):

$$\begin{aligned}\Pr(X \leq x) &= \Pr(F^{-1}(U) \leq x) \\ &= \Pr(U \leq F(x)) \\ &= F(x).\end{aligned}$$

So  $X = F^{-1}(U)$  is distributed according to  $F$ .

An important practical consequence of this observation is a process for sampling from an arbitrary CDF:

1. Sample from the uniform distribution to obtain a value of  $F(x)$ .
2. Apply  $F^{-1}$  to that sample to recover  $x = F^{-1}(F(x))$ , which is necessarily distributed according to  $F$ .

This process is called **inverse transform sampling**.

**Example 3.1.** The uniform distribution on  $[0, k]$  has CDF  $F(x) = x/k$  and inverse  $F^{-1}(y) = ky$ . Using inverse transform sampling, we can sample  $y$  from  $U[0, 1]$ , and then obtain as our sample  $F^{-1}(y) = F^{-1}(F(x)) = F^{-1}(x/k) = k(x/k) = x$ . From the transform, we know that this is equivalent to sampling  $x$  from  $U[0, k]$ .

**Example 3.2.** The exponential distribution has CDF  $F(x) = 1 - e^{-\lambda x}$ . We invert this CDF as follows:

$$\begin{aligned}F(x) &= 1 - e^{-\lambda x} \\ e^{-\lambda x} &= 1 - F(x) \\ -\lambda x &= \ln(1 - F(x)) \\ x &= -\ln(1 - F(x))/\lambda \\ F^{-1}(y) &= -\ln(1 - y)/\lambda.\end{aligned}$$

Now, using inverse transform sampling, we can sample from the exponential distribution by first sampling a value  $y$  from  $U[0, 1]$ , and then obtain as our sample  $F^{-1}(y) = F^{-1}(F(x)) = F^{-1}(1 - e^{-\lambda x}) = -\ln(1 - (1 - e^{-\lambda x}))/\lambda = \lambda x/\lambda = x$ . From the transform, we know that this is equivalent to sampling  $x$  from the exponential distribution.