CS145: Probability & Computing Lecture 7: Multiple Discrete Variables: Independence, Expectation



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Figure credits: Bertsekas & Tsitsiklis, **Introduction to Probability**, 2008 Pitman, **Probability**, 1999

CS145: Lecture 6 Outline

- Independent random variables
- Expectations of multiple discrete variables

Discrete Random Variables

 $\begin{array}{l} & \blacktriangleright \ \text{A random variable assigns values to outcomes of uncertain experiments} \\ & X:\Omega \to \mathbb{R} \qquad \qquad x = X(\omega) \in \mathbb{R} \ \text{ for } \omega \in \Omega \end{array}$



> The range of a random variable is the set of values with positive probability $\mathcal{X} = \{ x \in \mathbb{R} \mid X(\omega) = x \text{ for some } \omega \in \Omega, P(\omega) > 0 \}$

For a *discrete random variable*, the range is finite or countably infinite (we can map it to the integers). *Coming later: continuous random variables.*

Discrete Random Variables

 $\begin{array}{l} & \textbf{A random variable} \text{ assigns values to outcomes of uncertain experiments} \\ & X:\Omega \to \mathbb{R} \qquad \qquad x = X(\omega) \in \mathbb{R} \ \text{ for } \omega \in \Omega \end{array}$



> The probability mass function (PMF) or probability distribution of variable:

 $\sum p_X(x) = 1.$

 $x \in \mathcal{X}$

 $p_X(x) \ge 0,$

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega \mid X(\omega) = x\})$$

If range is finite, this is a vector of non-negative numbers that sums to one.

Joint Probability Distributions



In this example, N=2 and M=8, and the joint PMF is a 2x8 matrix.

- Consider two random variables X, Y. Suppose range of X is size N, range of Y is size M.
- > The *joint probability mass function* or *joint distribution* of two variables:

$$p_{XY}(x,y) = P(X = x \text{ and } Y = y)$$

$$p_{XY}(x,y) \ge 0, \qquad \sum_{x} \sum_{y} p_{XY}(x,y) = 1.$$

The joint distribution is uniquely specified by NM-1 numbers

Marginal Probability Distributions



The marginal distributions are defined by (N-1)+(M-1) numbers. Many joint distributions may have the same marginals.

The joint probability mass function or joint distribution of two variables:

$$p_{XY}(x,y) = P(X = x \text{ and } Y = y)$$

The range of each variable defines a partition of the sample space, so the *marginal distributions* can be computed from the joint distribution:

$$p_X(x) = P(X = x) = \sum_y p_{XY}(x, y) p_Y(y) = P(Y = y) = \sum_x p_{XY}(x, y)$$

Marginal Probability Distributions



Conditional Probability Distributions



> By the definition of conditional probability:

$$P(X = x \mid Y = y) = \frac{P(X = x \text{ and } Y = y)}{P(Y = y)}$$

> The conditional probability mass function is then:

$$p_{X|Y}(x \mid y) = P(X = x \mid Y = y) = \frac{p_{XY}(x, y)}{p_Y(y)} = \frac{p_{XY}(x, y)}{\sum_{x'} p_{XY}(x', y)}$$

Example: The Absent-Minded Prof



At office hours, a Professor gets 0, 1, or 2 questions with equal probability
Each question is answered correctly with probability ³/₄ (independently)

Several Random Variables $p_{XYZ}(x, y, z) = P(X = x \text{ and } Y = y \text{ and } Z = z)$



 $p_{XY}(x,y) \stackrel{\mathbf{x}}{=} \sum_{z \in \mathcal{Z}} p_{XYZ}(x,y,z) \qquad p_X(x) = \sum_{y \in \mathcal{Y}} p_{XY}(x,y) \qquad p_{XY|Z}(x,y \mid z) = \frac{p_{XYZ}(x,y,z)}{p_Z(z)}$

Marginal and conditional define new probability spaces.

May compute marginal and conditioned on any other set of variables.

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Reminder: Independence of Events

Independence of Two Events: $P(A \cap B) = P(A)P(B)$ This implies that $P(A \mid B) = P(A), P(B \mid A) = P(B).$

- > Observing B provides no information about whether A occurred
- Observing A provides no information about whether B occurred

Definition of Conditional Probabilities:

• **Definition:** Assuming $P(B) \neq 0$,

$$\mathbf{P}(A \mid B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)}$$

$$P(A | B)$$
 undefined if $P(B) = 0$



Independent Random Variables



> Equivalent conditions on conditional probabilities: $p_{X|Y}(x \mid y) = p_X(x) \text{ for all } p_Y(y) > 0$ $p_{Y|X}(y \mid x) = p_Y(y) \text{ for all } p_X(x) > 0$

Independent Random Variables



- For a given set of marginal distributions, there exists a unique joint distribution under which those variables are independent
- > Three random variables are independent if and only if

$$p_{XYZ}(x, y, z) = p_X(x)p_Y(y)p_Z(z)$$

Example: Independence



Verify that *X* and *Y* are *not* independent:

$$p_X(x) = p_Y(y) =$$

Conditional Independence



- Apply the same definition of independence for X and Y, but condition all probability distributions on some other variable Z
- Independence does not always imply conditional independence, and conditional independence does not always imply independence

Example: (Conditional) Independence



Verify that *X* and *Y* are *not* independent:

 $p_X(x) = p_Y(y) =$



But X and Y are *conditionally independent* given

 $Z = 1_{\{X \le 2, Y \ge 3\}}$ $p_{X|Z}(x \mid 1) =$ $p_{Y|Z}(y \mid 1) =$

Example:

Consider the following game: We roll a dice until we obtain an even number.

Define: X = the number of rolls in a game, Y= the value of the last roll (Y can be either 2, 4 or 6).

1) What is the distribution of (X,Y)?

2) What is the distribution of X?

3) What is the distribution of Y?

4) Are X and Y independent?

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Expectation

> The *expectation* or *expected value* of a discrete random variable is:

$$E[X] = \sum_{x \in \mathcal{X}} x p_X(x)$$

The expectation is a single number, not a random variable.
It encodes the "center of mass" of the probability distribution:

$$x_{\min} \le E[x] \le x_{\max} \qquad \begin{aligned} x_{\min} = \min\{x \mid x \in \mathcal{X}\} \\ x_{\max} = \max\{x \mid x \in \mathcal{X}\} \end{aligned}$$

> The expectation is an average or interpolation. It is possible that $p_X(E[x]) = 0$ for some random variables X.

Expected Values of Functions

 $\mathbf{V} = \mathbf{v}(\mathbf{V})$

Consider a non-random (deterministic) function of a random variable:

$$p_X(x) = P(X = x)$$
 $p_Y(x) = \sum_{\{x | g(x) = y\}} p_X(x)$

- > What is the expected value of random variable Y? E[Y] = E[g(X)]
- Correct approach #1:
- Correct approach #2:
- Incorrect approach:

$$E[Y] = \sum_{y} y p_{Y}(y)$$
$$E[Y] = E[g(X)] = \sum_{x} g(x) p_{X}(x)$$

 $g(E[X]) \neq E[g(X)] \quad \text{(except in special cases)}$

Examples

$$X = 1,2,3$$
 with probability 1/3,1/3,1/3
 $Y = X^2$

$$E[Y] = E[X^2] = \frac{1}{2} + \frac{4}{3} + \frac{9}{3} = 4\frac{2}{3}$$

$$(E[X])^2 = 2^2 = 4$$

$$X = -1, 0, +1$$
 with probabity $1/3, 1/3, 1/3$.
 $E[X^2] = \frac{2}{3}$, while $(E[X])^2 = 0$

Expectation of Multiple Variables

- > The *expectation* or *expected value* of a function of two discrete variables: $E[g(X,Y)] = \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} g(x,y) p_{XY}(x,y)$
- > A similar formula applies to functions of 3 or more variables
- Expectations of sums of functions are sums of expectations: $E[g(X) + h(Y)] = E[g(X)] + E[h(Y)] = \left[\sum_{x \in \mathcal{X}} g(x)p_X(x)\right] + \left[\sum_{y \in \mathcal{Y}} h(y)p_Y(y)\right]$
- > This is always true, whether or not X and Y are independent
- > Specializing to *linear functions*, this implies that:

$$E[aX + bY + c] = aE[X] + bE[Y] + c$$

Examples:

We role 2 die and get the sum of the first role plus the square of the second role

$$E[X + Y^{2}] = \sum_{x=1}^{6} \sum_{y=1}^{6} (x + y^{2}) Pr(X = x, Y = y)$$

=
$$\sum_{x=1}^{6} \sum_{y=1}^{6} x Pr(X = x, Y = y) + \sum_{x=1}^{6} \sum_{y=1}^{6} y^{2} Pr(X = x, Y = y)$$

=
$$\sum_{x=1}^{6} x Pr(X = x) + \sum_{x=1}^{6} y^{2} Pr(Y = y)$$

=
$$E[X] + E[Y^{2}]$$

Mean of Binomial Probability Distribution

- Suppose you flip *n* coins with bias *p*, count number of heads
- A binomial random variable X has parameters n, p:

$$p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$$

For binomial, expected values are expected counts of events:

$$E[X] = pn$$

Simple proof uses indicator variables X_i for whether each of n tosses is heads:

$$E[X_i] = p \cdot 1 + (1 - p) \cdot 0 = p = \Pr(X_i = 1).$$
$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = np.$$

$$\mathcal{X} = \{0, 1, 2, \dots, n\}$$



Binomial Mean: The Hard Way

$$\begin{split} \mathbf{E}[X] &= \sum_{j=0}^{n} j {n \choose j} p^{j} (1-p)^{n-j} \\ &= \sum_{j=0}^{n} j \frac{n!}{j!(n-j)!} p^{j} (1-p)^{n-j} \\ &= \sum_{j=1}^{n} \frac{n!}{(j-1)!(n-j)!} p^{j} (1-p)^{n-j} \\ &= np \sum_{j=1}^{n} \frac{(n-1)!}{(j-1)!((n-1)-(j-1))!} p^{j-1} (1-p)^{(n-1)-(j-1)} \\ &= np \sum_{k=0}^{n-1} \frac{(n-1)!}{k!((n-1)-k)!} p^{k} (1-p)^{(n-1)-k} \\ &= np \sum_{k=0}^{n-1} {n-1 \choose k} p^{k} (1-p)^{(n-1)-k} = np. \end{split}$$